

## Reliability Exponential Distribution (1+3) Cascade Model

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### Abstract

In this paper, a reliability formula was found for a cascade model containing four units (the first unit  $\mathcal{D}_1$  is basic and the other three units  $\mathcal{D}_2$ ,  $\mathcal{D}_3$ , and  $\mathcal{D}_4$  are redundant standby). It was assumed that the two variables strength and stress follow the exponential distribution of the variables. The estimation of the exponential distribution parameters was found using three estimation methods (Maximum likelihood, Percentile, and Least Squares), these methods were used to estimate the reliability of the model. MATLAB was used in the Monte Carlo simulation and mean square errors was used to compare the simulation results and find which estimation methods are the best for reliability estimation. The simulation results proved that the estimator ML is the best for estimating the model's reliability.

### 1. Introduction:

The interest in reliability has increased with the increasing dependence on industrial systems, which have become more complex, so it was necessary to pay attention to the life span of these systems and maintain their continued operation for a longer period. Using the mathematical formula  $R = (X \geq Y)$  [1-2], reliability can be calculated, where  $X$  stands for the strength and  $Y$  refers to the for stress, where the unit remains working if  $X$  is greater than  $Y$  and the unit fails if  $y$  becomes greater than  $X$ , so many researchers were interested in looking for the possibility of increasing reliability by increasing the strength of in this way, the durability of the new unit is multiplied by the strength attenuation  $m$  factor and its interface by Stress attenuation  $k$  [3-5].

In this paper, the general formula of reliability will be found for a model containing four units (one main unit and three standby units) where the stress r.v. and strength follow the exponential distribution and reliability estimation is found by methods (ML, Pr, LS), and the results were compared using MSE [6-8].

## 2. Model

In the cascade model (1+3) there are four life periods, and this depends on the sequence of the working unit in the order of the model, as this model consists of four units, namely  $\mathcal{D}_1$ ,  $\mathcal{D}_2$ ,  $\mathcal{D}_3$ , and  $\mathcal{D}_4$  basic units and three standby surplus units, so that when the active unit fails, it is replaced by one of the surplus units, according to order in the model. The reliability of the model is calculated in four stages, namely  $\mathcal{S}_1$ ,  $\mathcal{S}_2$ ,  $\mathcal{S}_3$  and  $\mathcal{S}_4$ , and in total, the total reliability of the model R is obtained as follows:

Suppose that random variables for strength and stress follow Exponential distribution where  $X \sim \text{Exp}(\beta)$  and  $Y \sim \text{Exp}(\eta)$  then the CDF of  $\text{Exp}(\beta)$  is:

$$F(x) = 1 - e^{-\beta x} \quad x > 0; \beta > 0 \quad \dots(1)$$

Now, the reliability will be found in each working case of the model and order:

The first stage: the basic unit  $\mathcal{D}_1$  is active and the rest of the units are surplus, the reliability calculation in this state is as follows:

$$\begin{aligned} \mathcal{S}_1 &= p[X_1 \geq Y_1] \\ &= \int_0^\infty [\bar{F}_{x_1}(y_1)] g(y_1) dy_1 \\ &= \int_0^\infty [e^{-\beta_1 y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\ \mathcal{S}_1 &= \int_0^\infty \eta_1 e^{-(\beta_1 + \eta_1) y_1} dy_1 \\ \mathcal{S}_1 &= \left[ \frac{\eta_1}{\beta_1 + \eta_1} \right] \quad \dots(2) \end{aligned}$$

The second stage: in this state, the basic unit  $\mathcal{D}_1$  fails and is replaced by unit  $\mathcal{D}_2$ , and the reliability is calculated as follows:

$$\mathcal{S}_2 = \text{pr}[X_1 < Y_1, X_2 \geq Y_2] = \text{pr}[X_1 < Y_1, mX_1 \geq kY_1]$$

$$\begin{aligned} \mathcal{S}_2 &= \int_0^\infty [F_{x_1}(y_1)] \left[ \bar{F}_{x_1} \left( \frac{k}{m} y_1 \right) \right] g(y_1) dy_1 \\ \mathcal{S}_2 &= \int_0^\infty [1 - e^{-\beta_1 y_1}] \left[ e^{-\beta_1 \left( \frac{k}{m} y_1 \right)} \right] \eta_1 e^{-\eta_1 y_1} dy_1 \end{aligned}$$

Let  $\frac{k}{m} = Z$ , then

$$\begin{aligned} \mathcal{S}_2 &= \int_0^\infty [e^{-\beta_1 Z y_1} - e^{-\beta_1 (1+Z) y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\ &= \int_0^\infty \eta_1 e^{-(\beta_1 Z + \eta_1) y_1} dy_1 - \int_0^\infty \eta_1 e^{-(\beta_1 (1+Z) + \eta_1) y_1} dy_1 \\ \mathcal{S}_2 &= \left( \frac{\eta_1}{\beta_1 Z + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1 (1+Z) + \eta_1} \right) \\ \mathcal{S}_2 &= \beta_1 \eta_1 \left[ \frac{1}{\beta_1 Z + \eta_1} \right] \left[ \frac{1}{\beta_1 (1+Z) + \eta_1} \right] \quad \dots(3) \end{aligned}$$

The third stage: when unit  $\mathcal{D}_2$  fails and is replaced by unit  $\mathcal{D}_3$ , the reliability can be calculated as follows:

$$\begin{aligned} \mathcal{S}_3 &= \text{pr}[X_1 < Y_1, X_2 < Y_2, X_3 \geq Y_3] \\ \mathcal{S}_3 &= \text{pr}[X_1 < Y_1, mX_1 < kY_1, mX_2 \geq kY_2] \\ \mathcal{S}_3 &= \text{pr}[X_1 < Y_1, mX_1 < kY_1, m^2 X_1 \geq k^2 Y_1] \\ \mathcal{S}_3 &= \int_0^\infty [F_{x_1}(y_1)] \left[ F_{x_1} \left( \frac{k}{m} y_1 \right) \right] \left[ \bar{F}_{x_1} \left( \frac{k^2}{m^2} y_1 \right) \right] g(y_1) dy_1 \\ \mathcal{S}_3 &= \int_0^\infty [1 - e^{-\beta_1 y_1}] [1 - e^{-\beta_1 Z y_1}] [e^{-\beta_1 Z^2 y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\ \mathcal{S}_3 &= \int_0^\infty [1 - e^{-\beta_1 y_1} - e^{-\beta_1 Z y_1} + e^{-\beta_1 (1+Z) y_1}] [e^{-\beta_1 Z^2 y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\ \mathcal{S}_3 &= \int_0^\infty [e^{-\beta_1 Z^2 y_1} - e^{-\beta_1 (1+Z^2) y_1} - e^{-\beta_1 (Z+Z^2) y_1} + e^{-\beta_1 (1+Z+Z^2) y_1}] \\ &\quad \cdot \eta_1 e^{-\eta_1 y_1} dy_1 \\ \mathcal{S}_3 &= \int_0^\infty \eta_1 e^{-(\beta_1 Z^2 + \eta_1) y_1} dy_1 - \int_0^\infty \eta_1 e^{-(\beta_1 (1+Z^2) + \eta_1) y_1} dy_1 \end{aligned}$$

$$\begin{aligned}
& - \int_0^\infty \eta_1 e^{-(\beta_1(Z+Z^2)+\eta_1)y_1} dy_1 + \int_0^\infty \eta_1 e^{-(\beta_1(1+Z+Z^2)+\eta_1)y_1} dy_1 \\
\mathcal{S}_3 &= \left( \frac{\eta_1}{\beta_1 Z^2 + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1(1+Z^2) + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1(Z+Z^2) + \eta_1} \right) + \left( \frac{\eta_1}{\beta_1(1+Z+Z^2) + \eta_1} \right) \\
\mathcal{S}_3 &= 2\eta_1 \beta_1^2 Z \left[ \frac{1}{(\beta_1 Z^2 + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z^2) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(Z+Z^2) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z+Z^2) + \eta_1)} \right] \quad \dots(4)
\end{aligned}$$

The fourth stage: when unit  $\mathcal{D}_3$  fails and is replaced by unit  $\mathcal{D}_4$ , reliability is calculated as follows:

$$\begin{aligned}
\mathcal{S}_4 &= \text{pr}[X_1 < Y_1, X_2 < Y_2, X_3 < Y_3, X_4 \geq Y_4] \\
\mathcal{S}_4 &= \text{pr}[X_1 < Y_1, mX_1 < kY_1, m^2X_1 < k^2Y_1, m^3X_1 \geq k^3Y_1] \\
\mathcal{S}_4 &= \int_0^\infty [F_{x_1}(y_1)] \left[ F_{x_1} \left( \frac{k}{m} y_1 \right) \right] \left[ F_{x_1} \left( \frac{k^2}{m^2} y_1 \right) \right] \left[ \bar{F}_{x_1} \left( \frac{k^3}{m^3} y_1 \right) \right] g(y_1) dy_1 \\
\mathcal{S}_4 &= \int_0^\infty [1 - e^{-\beta_1 y_1}] [1 - e^{-\beta_1 Z y_1}] [1 - e^{-\beta_1 Z^2 y_1}] [e^{-\beta_1 Z^3 y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\
\mathcal{S}_4 &= \int_0^\infty [1 - e^{-\beta_1 y_1} - e^{-\beta_1 Z y_1} + e^{-\beta_1(1+Z)y_1}] [1 - e^{-\beta_1 Z^2 y_1}] [e^{-\beta_1 Z^3 y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\
\mathcal{S}_4 &= \int_0^\infty [1 - e^{-\beta_1 y_1} - e^{-\beta_1 Z y_1} + e^{-\beta_1(1+Z)y_1} - e^{-\beta_1 Z^2 y_1} + e^{-\beta_1(1+Z^2)y_1} + e^{-\beta_1(Z+Z^2)y_1} \\
& \quad - e^{-\beta_1(1+Z+Z^2)y_1}] [e^{-\beta_1 Z^3 y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\
&= \int_0^\infty [e^{-\beta_1 Z^3 y_1} - e^{-\beta_1(1+Z^3)y_1} - e^{-\beta_1(Z+Z^3)y_1} + e^{-\beta_1(1+Z+Z^3)y_1} - e^{-\beta_1(Z^2+Z^3)y_1} \\
& \quad + e^{-\beta_1(1+Z^2+Z^3)y_1} + e^{-\beta_1(Z+Z^2+Z^3)y_1} - e^{-\beta_1(1+Z+Z^2+Z^3)y_1}] \eta_1 e^{-\eta_1 y_1} dy_1 \\
\mathcal{S}_4 &= \int_0^\infty \eta_1 e^{-(\beta_1 Z^3 + \eta_1 \rho_1) y_1} dy_1 - \int_0^\infty \eta_1 e^{-(\beta_1(1+Z^3) + \eta_1 \rho_1) y_1} dy_1 \\
& \quad - \int_0^\infty \eta_1 e^{-(\beta_1(Z+Z^3) + \eta_1 \rho_1) y_1} dy_1 + \int_0^\infty \eta_1 e^{-(\beta_1(1+Z+Z^3) + \eta_1 \rho_1) y_1} dy_1 \\
& \quad - \int_0^\infty \eta_1 e^{-(\beta_1(Z^2+Z^3) + \eta_1 \rho_1) y_1} dy_1 + \int_0^\infty \eta_1 e^{-(\beta_1(1+Z^2+Z^3) + \eta_1 \rho_1) y_1} dy_1 \\
& \quad + \int_0^\infty \eta_1 e^{-(\beta_1(Z+Z^2+Z^3) + \eta_1 \rho_1) y_1} dy_1 - \int_0^\infty \eta_1 e^{-(\beta_1(1+Z+Z^2+Z^3) + \eta_1 \rho_1) y_1} dy_1 \\
\mathcal{S}_4 &= \left( \frac{\eta_1}{\beta_1 Z^3 + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1(1+Z^3) + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1(Z+Z^3) + \eta_1} \right) + \left( \frac{\eta_1}{\beta_1(1+Z+Z^3) + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1(Z^2+Z^3) + \eta_1} \right) \\
& \quad + \left( \frac{\eta_1}{\beta_1(1+Z^2+Z^3) + \eta_1} \right) + \left( \frac{\eta_1}{\beta_1(Z+Z^2+Z^3) + \eta_1} \right) - \left( \frac{\eta_1}{\beta_1(1+Z+Z^2+Z^3) + \eta_1} \right) \\
\mathcal{S}_4 &= 2\eta_1 \beta_1^2 Z \left[ \frac{1}{(\beta_1 Z^3 + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(Z+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z+Z^3) + \eta_1)} \right] \\
& \quad - 2\eta_1 \beta_1^2 Z \left[ \frac{1}{(\beta_1(Z^2+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z^2+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(Z+Z^2+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z+Z^2+Z^3) + \eta_1)} \right] \quad \dots(5)
\end{aligned}$$

Finally, by combining the above four states, the general formula of reliability of the (3+1) cascade model is obtained:

$$\begin{aligned}
R &= \left[ \frac{\eta_1}{\beta_1 + \eta_1} \right] + \beta_1 \eta_1 \left[ \frac{1}{\beta_1 Z + \eta_1} \right] \left[ \frac{1}{\beta_1(1+Z) + \eta_1} \right] + 2\eta_1 \beta_1^2 Z \left[ \frac{1}{(\beta_1 Z^2 + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z^2) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(Z+Z^2) + \eta_1)} \right] \\
& \quad \cdot \left[ \frac{1}{(\beta_1(1+Z+Z^2) + \eta_1)} \right] + 2\eta_1 \beta_1^2 Z \left[ \frac{1}{(\beta_1 Z^3 + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(Z+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z+Z^3) + \eta_1)} \right] \\
& \quad - 2\eta_1 \beta_1^2 Z \left[ \frac{1}{(\beta_1(Z^2+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z^2+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(Z+Z^2+Z^3) + \eta_1)} \right] \left[ \frac{1}{(\beta_1(1+Z+Z^2+Z^3) + \eta_1)} \right] \quad \dots(6)
\end{aligned}$$

### 3. Estimation:

The exponential distribution parameters are estimated using three different estimation methods as follows:

#### 3-1 Maximum likelihood

Let the random sample  $x_1, x_2, \dots, x_n$  from  $\text{Exp}(\beta)$ , the likelihood function "L", is [9-11]:

$$\begin{aligned}
L(x_1, x_2, \dots, x_n, \beta) &= f(x_1; \beta) f(x_2; \beta) \dots f(x_n; \beta) \\
&= \prod_{i=1}^n f(x_i; \beta) \quad \dots(7)
\end{aligned}$$

So the equation 7 becomes:

$$L(x_1, x_2, \dots, x_n; \beta) = \beta^n \prod_{i=1}^n e^{-\sum_{i=1}^n \beta x} \quad \dots(8)$$

Equation 8 becomes after taking its logarithm and deriving it as written below

$$\ln L = n \ln \beta - \beta \sum_{i=1}^n x$$

$$\frac{\partial \ln L}{\partial \beta} = \frac{n}{\beta} - \sum_{i=1}^n x$$

$$\frac{n}{\beta} - \sum_{i=1}^n x = 0$$

So the maximum likelihood estimator of  $\beta$ :

$$\hat{\beta}_{ML} = \frac{n}{\sum_{i=1}^n x} \quad \dots(9)$$

And

$$\hat{\eta}_{ML} = \frac{m}{\sum_{j=1}^m y} \quad \dots(10)$$

### 3-2 Percentile Method

Let  $x_i$ ;  $i = 1, 2, \dots, n$  of  $\text{Exp}(\beta)$ , by using equation 1, [12]:

$$\ln(1 - F(x_{(i)})) = -\beta x_{(i)}$$

$$x_{(i)} = \left( \frac{-\ln(1 - F(x_{(i)}))}{\beta} \right) \quad \dots(11)$$

$$\ln[(1 - P_i)^{-1}] = \beta x_{(i)} \quad \dots(12)$$

Where  $P_i$  is plotting position and  $P_i = \frac{i}{n+1}$ ;  $i = 1, 2, \dots, n$

$$x_{(i)} = \left( \frac{-\ln(1 - P_i)}{\beta} \right) \quad \dots(13)$$

The minimizing equation is:

$$\sum_{i=1}^n [x_{(i)} - F(x_{(i)})]^2 \quad \dots(14)$$

Substitution 13 in 14, then:

$$\sum_{i=1}^n \left[ x_{(i)} - \left( \frac{-\ln(1 - P_i)}{\beta} \right) \right]^2 \quad \dots(15)$$

By deriving the equation 15:

$$\sum_{i=1}^n 2 \left[ (x_{(i)} - \beta^{-1}(-\ln(1 - P_i))) (\beta^{-2})(-\ln(1 - P_i)) \right] = 0$$

So that:

$$\hat{\beta}_{Pr} = \left[ \frac{\sum_{i=1}^n (-\ln(1 - P_i))^2}{\sum_{i=1}^n (x_{(i)}) (-\ln(1 - P_i))} \right] \quad \dots(16)$$

and

$$\hat{\eta}_{Pr} = \left[ \frac{\sum_{j=1}^m (-\ln(1 - P_j))^2}{\sum_{j=1}^m (y_{(j)}) (-\ln(1 - P_j))} \right] \quad \dots(17)$$

### 3-1 Least Squares Method

Let  $x_1, x_2, \dots, x_n$  with  $\text{Exp}(\beta)$ , the equation is used as follow, [13-16]:

$$\begin{aligned} Q &= \sum_{i=1}^r (\hat{F}(x_{(i)}) - F(x_{(i)}))^2 \\ &= \sum_{i=1}^n (\hat{F}(x_{(i)}) - (1 - e^{-\beta x}))^2 \end{aligned} \quad \dots(18)$$

So that:

$$-\ln(1 - F(x_{(i)})) = \theta x_{(i)}^\sigma \quad \dots(19)$$

$$\text{Let } q_{(i)} = -\ln(1 - \hat{F}(x_{(i)})) = -\ln(1 - P_i)$$

Then Equation 19 becomes:

$$S(2, \lambda) = \sum_{i=1}^n (q_{(i)} - \beta x_{(i)})^2 \quad \dots(20)$$

So that:

$$\frac{\partial S(\beta)}{\partial \beta} = \sum_{i=1}^n 2(q_{(i)} - \beta x_{(i)}) (-x_{(i)}) = 0$$

$$-\sum_{i=1}^n q_{(i)} x_{(i)} + \beta \sum_{i=1}^n x_{(i)}^2 = 0$$

Then  $\hat{\beta}_{LS}$  :

$$\hat{\beta}_{LS} = \frac{\sum_{i=1}^n q_{(i)} x_{(i)}}{\sum_{i=1}^n x_{(i)}^2} \quad \dots(21)$$

and

$$\hat{\eta}_{LS} = \frac{\sum_{j=1}^m q_{(j)} y_{(j)}}{\sum_{j=1}^m y_{(j)}^2} \quad \dots(22)$$

#### 4. The simulation

Using the estimates obtained, a simulation is made to compare results of these estimates using MSE. Different small, medium, and large sample sizes were used, and different parameter values. Six experiments were conducted for various parameter values to compare the estimation results. Experiments made were founded on run size  $K=5000$ . Sample sizes are  $(n, m) = (20, 20), (40, 40),$  and  $(80, 80)$ , where  $\mathcal{N}_1 = (20, 20), \mathcal{N}_2 = (40, 40),$  and  $\mathcal{N}_3 = (80, 80)$ . The values of sex experiments are included in Table 1:

**Table 1:** The experiment values.

experiment	$k$	$m$	$\beta_1$	$\eta_1$	$R$
1	1.80	0.30	0.7	0.7	0.5179
2	1.20	0.80	0.7	0.7	0.6287
3	1.80	0.30	0.9	0.7	0.4523
4	1.20	0.80	0.9	0.7	0.5524
5	1.80	0.30	0.7	1.5	0.7106
6	1.20	0.80	0.7	1.5	0.8209

The results of the simulation for experiments:

**Table 2:** Simulation of experiment (1)

S.S.	Create	ML	Pr	LS	Favorite
$\mathcal{N}_1$	Mean	0.5192	0.5194	0.5195	
	MSE	0.0068	0.0076	0.0082	ML
$\mathcal{N}_2$	Mean	0.5183	0.5179	0.5177	
	MSE	0.0034	0.0040	0.0043	ML
$\mathcal{N}_3$	Mean	0.5187	0.5190	0.5191	
	MSE	0.0017	0.0020	0.0022	ML

**Table 3:** Simulation of experiment (2)

S.S.	Create	ML	Pr	LS	Favorite
$\mathcal{N}_1$	Mean	0.6248	0.6262	0.5950	
	MSE	0.0083	0.0092	0.0099	ML
$\mathcal{N}_2$	Mean	0.6282	0.6294	0.5987	ML

$\mathcal{N}_3$	MSE	0.0041	0.0048	0.0055	
	Mean	0.6290	0.6291	0.5988	
	MSE	0.0022	0.0025	0.0033	ML

**Table 4:** Simulation of experiment (3)

S.S.	Create	ML	Pr	LS	Favorite
$\mathcal{N}_1$	Mean	0.5190	0.5193	0.5193	
	MSE	0.0111	0.0119	0.0125	ML
$\mathcal{N}_2$	Mean	0.5172	0.5172	0.5175	
	MSE	0.0077	0.0082	0.0086	ML
$\mathcal{N}_3$	Mean	0.5183	0.5186	0.5188	
	MSE	0.0062	0.0065	0.0066	ML

**Table 5:** Simulation of experiment (4)

S.S.	Create	ML	Pr	LS	Favorite
$\mathcal{N}_1$	Mean	0.6222	0.6235	0.5991	
	MSE	0.0110	0.0142	0.0121	ML
$\mathcal{N}_2$	Mean	0.6250	0.6254	0.6014	
	MSE	0.0095	0.0101	0.0171	ML
$\mathcal{N}_3$	Mean	0.6247	0.6250	0.6016	
	MSE	0.0074	0.0078	0.0089	ML

**Table 6:** Simulation of experiment (5)

S.S.	Create	ML	Pr	LS	Favorite
$\mathcal{N}_1$	Mean	0.5174	0.5172	0.5171	
	MSE	0.0443	0.0452	0.0457	ML
$\mathcal{N}_2$	Mean	0.5189	0.5191	0.5192	
	MSE	0.0402	0.0406	0.0409	ML
$\mathcal{N}_3$	Mean	0.5185	0.5182	0.5182	
	MSE	0.0386	0.0390	0.0392	ML

**Table 7:** Simulation of experiment (6)

S.S.	Create	ML	Pr	LS	Favorite
$\mathcal{N}_1$	Mean	0.8209	0.6267	0.5953	
	MSE	0.0466	0.0470	0.0597	ML
$\mathcal{N}_2$	Mean	0.6274	0.6279	0.5969	
	MSE	0.0417	0.0422	0.0549	ML
$\mathcal{N}_3$	Mean	0.6293	0.6298	0.5994	
	MSE	0.0389	0.0390	0.0514	ML

## 5. Conclusions

After conducting the six experiments, it was concluded that the parameter values affect the reliability values of the model. When the value of the parameter  $\beta$  increases, the reliability of the model decreases, while the value of the parameter increases with the value of the parameter  $\eta$ , and the reliability value decreases with an increase of  $\frac{k}{m}$ . This is evident when comparing the reliability values in Table 1.

After conducting simulations for the six experiments and for different sample sizes, it was concluded that estimator ML is the best for estimating the model's reliability and for all the results, as is clear in Tables 2 to 7.

## References

1. Rahman, J. ul, Mohyuddin, M. R., Anjum, N., & Butt, R. (2016). Modelling of two Interconnected Spring Carts and Minimization of Energy. *DJ Journal of Engineering and Applied Mathematics*, 2(1), 7–11.
2. Siju, K. C., & Kumar, M. (2016). Reliability analysis of time dependent stress-strength model with random cycle times. *Perspectives in Science*, 8, 654–657.
3. Khaleel, A. H. Ahmed R Khlefha1,(2021).Reliability of (1+1) Cascade Model for Weibull Distribution. *Highlights in Science*, 1, 1-4.
4. Patowary, A. N., Hazarika, J., & Sriwastav, G. L. (2018). Reliability estimation of multi-component cascade system through Monte-Carlo simulation. *International Journal of Systems Assurance Engineering and Management*, 9(6), 1279–1286.
5. Khaleel, A. H. ,(2024).Reliability for Generalized Rayleigh of 1 Strength - 4 Stresses. *JURNAL ILMIAH MATEMATIKA, SAINS DAN TEKNOLOGI*, 12(1), 26-31.
6. Jebur, I. G., Kalaf, B. A., & Salman, A. N. (2020). On Bayesian Estimation of System Reliability in Stress - Strength Model Based on Generalized Inverse Rayleigh Distribution. *IOP Conference Series: Materials Science and Engineering*, 871(1).
7. Hassan, A. S., Nagy, H. F., Muhammed, H. Z., & Saad, M. S. (2020). Estimation of multicomponent stress-strength reliability following Weibull distribution based on upper record values. *Journal of Taibah University for Science*, 14(1), 244–253.
8. Khaleel, A. H. ,(2024).Reliability Function of (1 Strength and 4 Stresses) for Rayleigh distribution. *Basrah Journal of Sciences*, 42(1), 13-22.
9. Kanaparathi, R., Palakurthi, J., & Narayana, L. (2020). Cascade reliability of stress-strength system for the new Rayleigh-Pareto Distribution, 5(2), 131–139.
10. Ashok, P., Devi, M. T., & Maheswari, T. S. U. (2019). Reliability of a cascade system of type  $(X<Y<Z)$  for Pareto and Weibull distributions. *AIP Conference Proceedings*, 2112(June).
11. Chaturvedi, A., & Malhotra, A. (2020). On Estimation of Stress-Strength Reliability Using Lower Record Values from Proportional Reversed Hazard Family. *American Journal of Mathematical and Management Sciences*, 39(3), 234–251.
12. A. H. Khaleel,(2021). Reliability of one Strength-four Stresses for Lomax Distribution. *J. Phys. Conf. Ser.*, 1879(3),1-10.
13. Khan, A. H., & Jan, T. R. (2014). Estimation of multi component systems reliability in stress-strength models. *Journal of Modern Applied Statistical Methods*, 13(2), 389–398.
14. Patowary, A. N., Hazarika, J., & Sriwastav, G. L. (2018). Reliability estimation of multi-component cascade system through Monte-Carlo simulation. *International Journal of Systems Assurance Engineering and Management*, 9(6), 1279–1286.

15. Salman, A. N., & Hamad, A. M. (2019). On estimation of the stress-Strength reliability based on lomax distribution. IOP Conference Series: Materials Science and Engineering, 571(1).
  16. Khaleel, A. H. ,(2024). THE FRECHET RELIABILITY FOR (2+2) CASCADE MODEL. African Journal of Mathematics and Statistics Studies, 7(1), 50-63.
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## معدلية التوزيع الاسي لنموذج كاسكاد (3+1)

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### الخلاصة:

في هذه البحث، تم إيجاد صيغة المعدلية لنموذج متتالي يحتوي على أربع وحدات (الوحدة الأولى الاساسية  $D_1$  والوحدات الثلاث الأخرى  $D_2$  و  $D_3$  و  $D_4$  احتياطية زائدة). تم افتراض أن متغيري المتانة والإجهاد يتبعان التوزيع الأسّي. تم إيجاد تقدير معالم التوزيع الأسّي باستخدام ثلاث طرق تقدير هي (الإمكان الاعظم، النسبة المؤوية، والمربعات الصغرى) ومن ثم تم تقدير موثوقية النموذج بهذه الطرق. تم إجراء محاكاة مونت كارلو باستخدام برنامج MATLAB لمقارنة نتائج طرق التقدير ومعرفة أيهما أفضل لتقدير موثوقية النموذج باستخدام متوسط مربع الخطأ وقد وجد أن مقدر الإمكان الأعظم هو الأفضل لتقدير معدلية النموذج.

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الكلمة المفتاحية: الوحدة، مونت كارل، النسبة المؤوية، والمربعات الصغرى، التوزيع الاسي

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